

Statistical Inference based on Inverse Data Generating Equation

(Generalized Fiducial Inference)

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- **Oxford English Dictionary**
 - **adjective** TECHNICAL (of a point or line) used as a fixed basis of comparison.
 - ORIGIN from Latin fiducia ‘trust, confidence’
- **Merriam-Webster dictionary**
 1. taken as standard of reference *a fiducial mark*
 2. founded on faith or trust
 3. having the nature of a trust : FIDUCIARY

Brief history of fiducial inference

- Fisher (1930) introduced the idea of fiducial probability and inference in an attempt to overcome what he saw as a serious deficiency of the Bayesian approach to inference – use of a prior distribution when no prior information was available.

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- Fisher (1935) further elaborated on this idea. E.g., to eliminate nuisance parameters he suggested substituting their fiducial distribution. As an example he considered the inference for the difference of two normal means – “Behrens-Fisher problem” .

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- Bernard (1995) pivotal based methods.
- Weerahandi (1989, 1993) generalized inference.
- Hannig, Iyer, Patterson (2006) generalized inference is closely related to fiducial inference & theoretical properties.

Related Modern Research

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- Confidence Distributions; [Xie, Singh & Strawderman \(2011\)](#), [Schweder & Hjort \(2002\)](#) The idea is to use a frequentist procedure (e.g., one sided CI for all possible confidence levels α) to define a distribution on the parameter space.

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- Objective Bayesian inference; choice of $\pi(\theta)$ when we have no prior info, e.g., reference prior [Berger, Bernardo & Sun \(2009\)](#).
 - With improper reference prior one needs to prove that the posterior is a proper distribution on an individual basis.

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 - Each statistical problem requires its own solution and the quality of the solution is judged by repeated sampling performance (Cournot's principle).

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- It is assumed that the value theta $\theta \in \Theta$ was generated using some known distribution $\pi(\theta)$, prior, and we have only single, fully known distribution $P_\theta \cdot \pi(\theta)$.

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- There is only one solution for each statistical problem. The remaining problem specific issue is to find the solution computationally and to select the right model + prior.

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- Philosophical interpretation of fiducial probability is obscure.
- We use fiducial distribution to propose statistical methods (e.g., confidence Intervals) and then evaluate the methods using repeated sampling performance.
- The fiducial distribution is usually not a posterior with respect to any (data independent) prior (Grundy, 1956).

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 - We attempt to strip down all layers of additional structure.
 - Our definition does not produce a "unique fiducial distribution".
Regardless, the fiducial distribution is always proper.
- We proved some asymptotic theorems justifying this method of deriving inference procedures. Simulations usually show very good frequentist performance.

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- **Likelihood** is the function $f(\mathbf{x}, \xi)$, where ξ is variable and \mathbf{x} is fixed.

Comparison to MLE

- Consider the **data generating** (structural) equation

$$\mathbf{X} = \mathbf{T}(U, \xi),$$

- U is a random variable/vector with known distribution
- ξ is a **fixed** parameter.
- The **distribution** of the data \mathbf{X} is implied from U via the structural equation. I.e., one can generate \mathbf{X} by generating U and plugging it into the structural equation.

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- If the solution does not exist, discard this value of U^* , i.e., condition the distribution of U on the fact that the solution exists.

Example – binomial

- Let X_1, \dots, X_n be i.i.d. Bernoulli(p). Therefore

$$X_i = I_{[0,p)}(U_i), \quad i = 1, \dots, n,$$

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- Define the inverse image of T

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- We need to select a point inside the interval. We recommend selecting each edge with equal probability.

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- We can simulate this distribution using $\mathcal{R}_\mu = 10 - Z^*$, where $Z^* \sim N(0, 1)$ independent of Z .

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- We have non-uniqueness due to Borel paradox.

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 - The choice among multiple solutions:
 - Arises if the inverse image $Q(\mathbf{x}, U^*)$ has more than one element but disappears asymptotically for parametric problems.
 - The conditioning on the fact that solution exist:
 - Arises if $P\{Q(\mathbf{x}, U^*) \neq \emptyset\} = 0$ – Borel paradox.
 - “Resolved by fat data”.

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- We derive generalized fiducial distribution directly for discretized data or take a limit as the discretization refines.

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$$\arg \min_{\xi} \|\mathbf{x} - T(\mathbf{U}^*, \xi)\| \mid \left\{ \min_{\xi} \|\mathbf{x} - T(\mathbf{U}^*, \xi)\| < \varepsilon \right\} \quad (1)$$

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- Similar to the idea of ABC; generating from prior replaced by **min**.

Theoretical results

- If using $\|\cdot\|_\infty$ and smooth T the limiting conditional distribution (1) has density (Hannig, 2012)

$$r(\xi|\mathbf{x}) = \frac{f_{\mathbf{X}}(\mathbf{x}|\xi)J(\mathbf{x}, \xi)}{\int_{\Xi} f_{\mathbf{X}}(\mathbf{x}|\xi')J(\mathbf{x}, \xi') d\xi'},$$

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where $J(\mathbf{x}, \xi) = \sum_{\mathbf{i}=(i_1, \dots, i_p)} \left| \det \left(\frac{d}{d\xi} \mathbf{T}(\mathbf{u}, \xi) \Big|_{\mathbf{u}=\mathbf{T}^{-1}(\mathbf{x}, \xi)} \right)_{\mathbf{i}} \right|$

and $(A)_{\mathbf{i}}$ is the $p \times p$ matrix comprising of the i_1, \dots, i_p th row of the $n \times p$ matrix A .

Comments

• Let $X_i = F^{-1}(\xi, U_i)$ be cont. with density $f(x|\xi)$.

• Then
$$J(\mathbf{x}, \xi) = \sum_{\mathbf{i}=(i_1, \dots, i_p)} \frac{|\det(\frac{d}{d\xi} \mathbf{F}(\mathbf{x}_i, \xi))|}{\prod_i f(x_i, \xi)}$$

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• Often $\binom{n}{p}^{-1} J(\mathbf{x}, \xi) \rightarrow E_{\xi_0} \frac{|\det(\frac{d}{d\xi} \mathbf{F}(\mathbf{X}_i, \xi))|}{\prod_i f(x_i, \xi)}$ providing an empirical Bayes interpretation.

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- Confidence intervals based on generalized fiducial distribution are often correct asymptotically because of “Bernstein-von Mises” theorem for fiducial distributions Hannig (2009, 2012), Sonderegger & Hannig (2012).

Theoretical result for discretized data

- Assume structural equation $X_i = F^{-1}(U_i, \xi)$
 - ξ is p dimensional and U_i are i.i.d. $U(0, 1)$.
 - $F(x, \xi)$ is continuously differentiable in ξ for all x
 - $(F(x_1, \xi), \dots, F(x_p, \xi)) = (u_1, \dots, u_p)$, taken as a function of ξ is one to one for each \mathbf{x} .

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 - $(F(x_1, \xi), \dots, F(x_p, \xi)) = (u_1, \dots, u_p)$, taken as a function of ξ is one to one for each \mathbf{x} .
- Data were discretized to a fixed partition $(-\infty, a_1], (a_1, a_2], \dots, (a_k, \infty)$.
 - $P(X \in (a_j, a_{j+1}]) > 0$ for all j .
 - For all $\mathbf{j} \subset \{1, \dots, k\}$, the Jacobian $\det \left(\frac{\mathbf{d}F(\mathbf{a}_{\mathbf{j}}, \xi_0)}{\mathbf{d}\xi} \right) \neq 0$.

Theoretical result for discretized data

- Assume structural equation $X_i = F^{-1}(U_i, \xi)$
 - ξ is p dimensional and U_i are i.i.d. $U(0, 1)$.
 - $F(x, \xi)$ is continuously differentiable in ξ for all x
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Theorem (Hannig (2012)). *Confidence sets based on the generalized fiducial distribution will have asymptotically correct coverage as number of data points goes to infinity and resolution remains fixed.*

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- [Wandler & Hannig \(2011, 2012\)](#) shows consistency for various multivariate normal model.

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- For simplicity assume, each laboratory reports a confidence interval based on a T distribution, measuring the same object.
- Goal is to combine the intervals in a way that down ways potential outliers. Outright dropping of odd results is politically not feasible.

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$$r(\mu) \propto \sum_{j \in I} C(\mathcal{M}_j) \sum_{i \in \mathcal{M}_j} \left\{ \frac{1}{n_i} + \frac{(\mu - \bar{x}_i)^2}{(n_i - 1)s_i^2} \right\}^{-1/2} \prod_{i \in \mathcal{M}_j} \left\{ 1 + \frac{n_i(\mu - \bar{x}_i)^2}{(n_i - 1)s_i^2} \right\}^{-(n_i - 1)/2} \\ \times e^{-(k - |\mathcal{M}_i|) \log(SSE)/2}$$

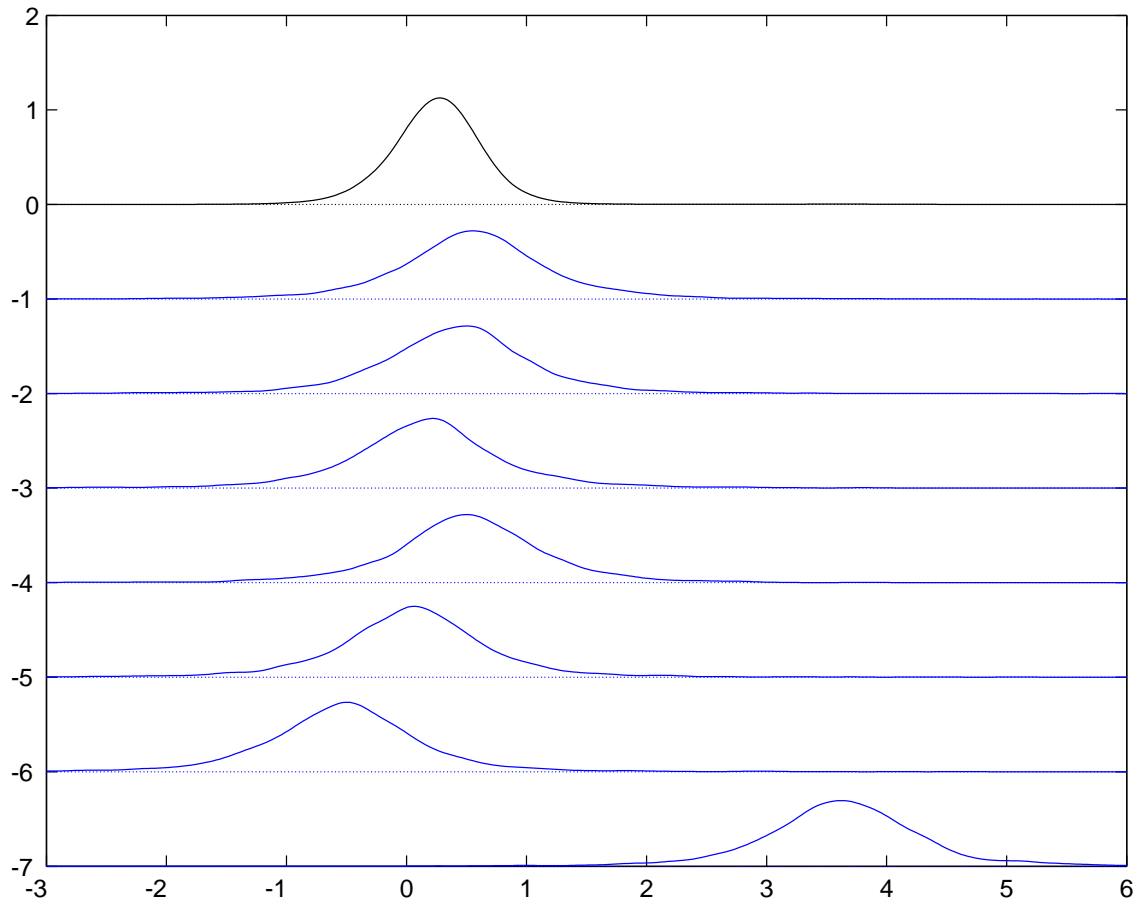
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- Preliminary simulation using importance sampling shows somewhat conservative performance.

Key comparison - example



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- There seems to be no (non-Bayesian) unified approach producing good quality confidence sets. Most procedures in the literature are designed to solve special cases [Burdick, Graybill, Wang](#) or use insufficient statistics [Khuri, Mathews and Sinha](#).
- We will propose a procedure that produces confidence sets for large class of linear mixed models. Additionally it allows for discretized data.

Linear Mixed Model

- Consider a structural equation

$$\mathbf{Y} = \mathbf{X}\beta + \sum_{i=1}^k \sigma_i \sum_{j=1}^{l_k} V_{i,j} Z_{i,j}$$

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- k number of random effects, l_k number of levels per effect,
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- Contains a wide variety of linear mixed models.

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One way random effects model

- $X = 1, k = 2, l_1$ number of levels for random effect, $l_2 = n$, $V_{1,i}$ indicates which observations are in group i , $V_{2,\cdot} = I$
- m overall mean, σ_1^2 random effect variance, σ_2^2 error variance

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- Need an efficient algorithm for generating such \mathbf{Z} .
- Possibilities include
 - Gibbs sampler – does not mix well if there is too much precision.
 - Simulated tampering – works but slow
 - We proposed a particular implementation of Sequential Monte Carlo algorithm
 - works well if the number of parameters is reasonable (< 10).

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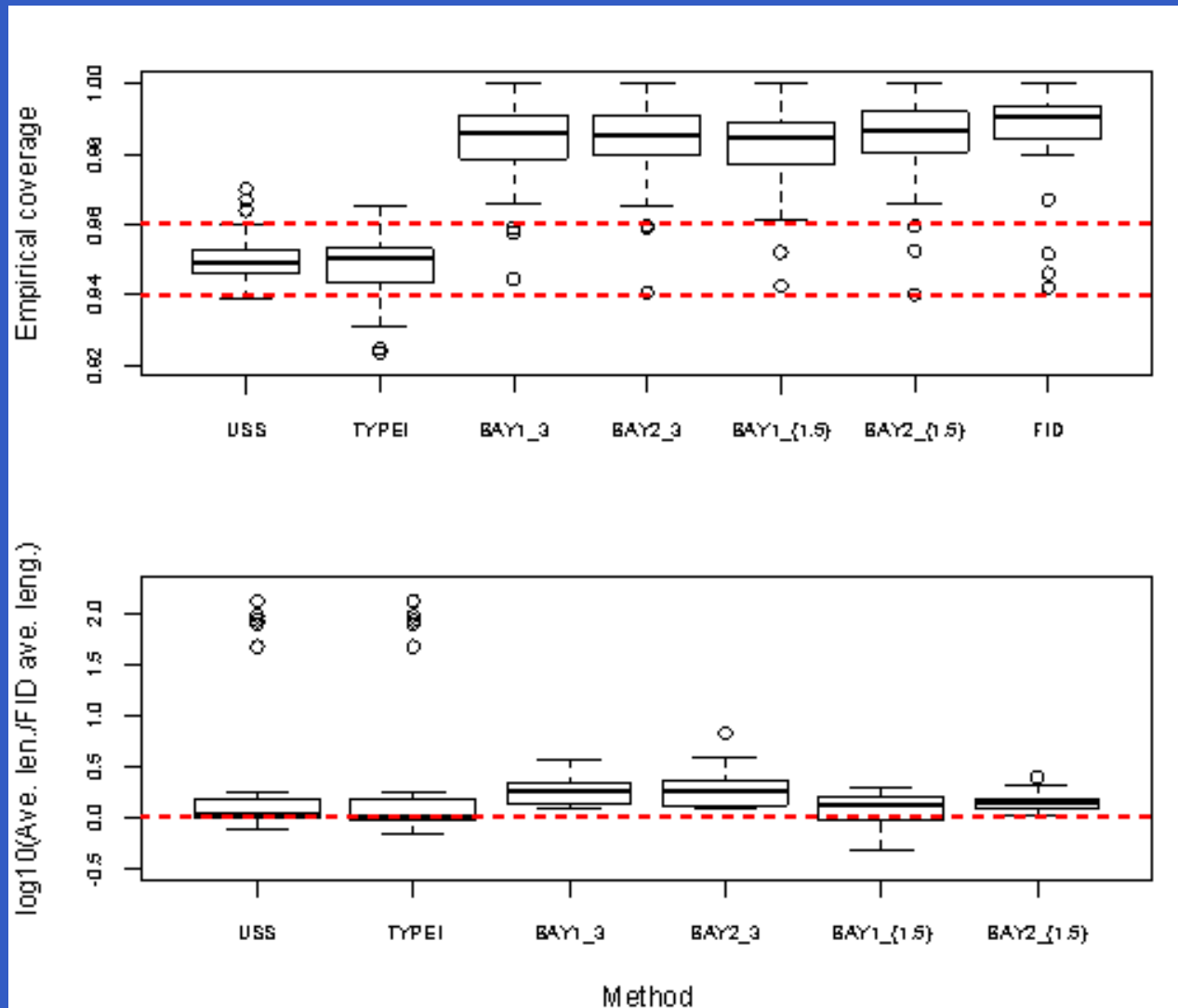
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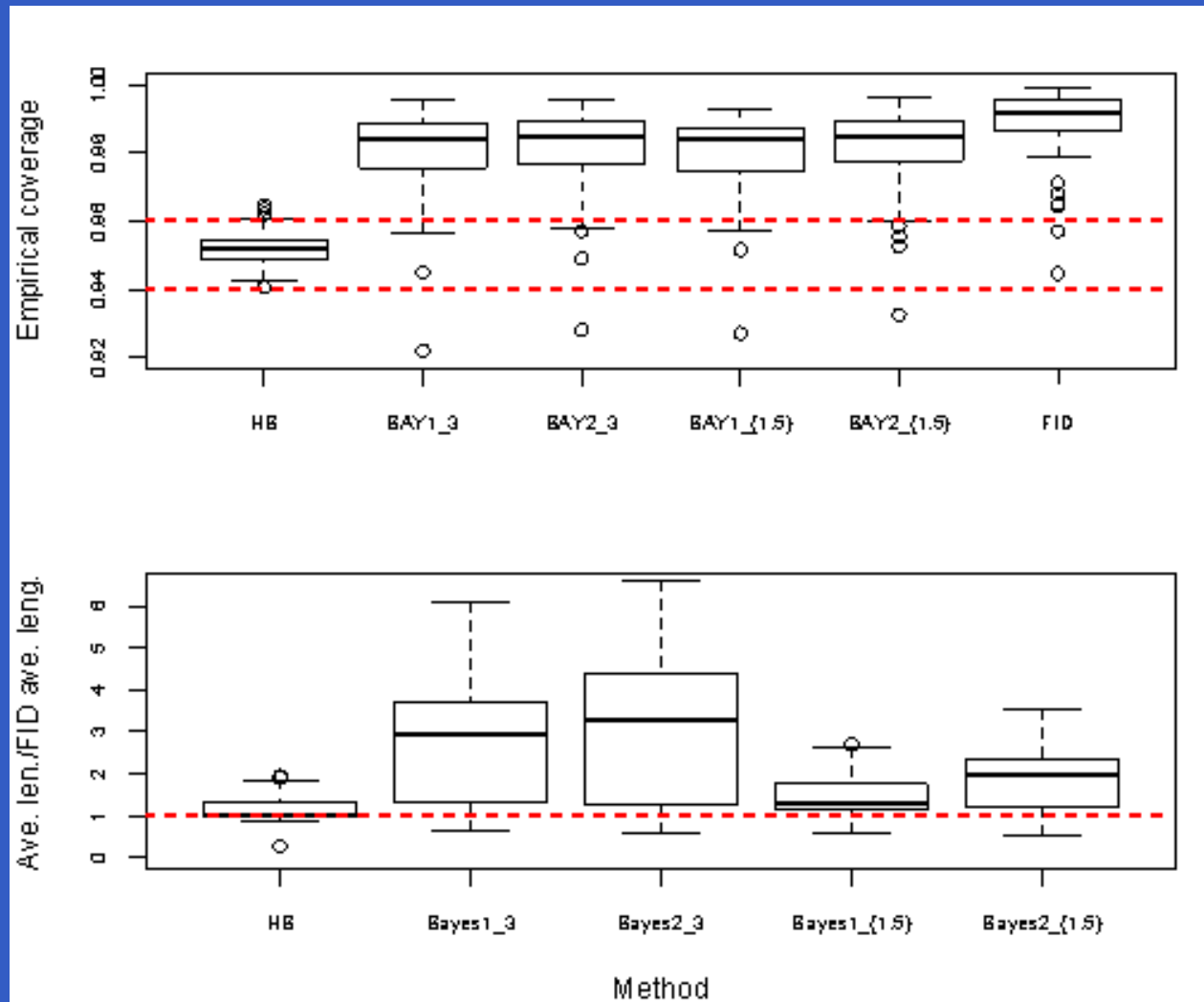
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(μ fixed, α , β , $(\alpha\beta)$, and ϵ are independent and \sim Normal)
- We considered a number of models with various levels of imbalance and values of parameters.

95% CI for random effects (nested)



95% CI for random effects (crossed)



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- Ultra-highdimensional Regression Model (How to properly introduce a penalty?)

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- Many simulation studies show that generalized fiducial solutions have very good small sample properties.
- Current popularity of generalized inference in some applied circles suggests that if computers were available 70 years ago, fiducial inference might not have been rejected.

Quotes

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